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International Price Linkage within a Region: The Case of East Asia

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Abstract

This paper investigates how "prices" in East Asian economies were correlated with those in Japan and with those in the United States. The analysis is particularly noteworthy because East Asian economies are geographically close to Japan but their currencies have been more tied to the US dollar. In the paper, we analyze two different types of "prices": the overall price levels in terms of the same currency and the relative prices among different commodities. We demonstrate that the overall price levels in East Asian economies were closely related to those in the United States. However, we also show that the relative prices in East Asian economies, especially those in Taiwan and Korea, were closely correlated with those in Japan. The results are in marked contrast with the price correlation in other regions.

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1. Introduction

The failure of the law of one price in international trade has been widely documented. One well-understood explanation for this failure is the Balassa (1964) - Samuelson (1964) effect which stresses the differential productivity growth between tradable and non-tradable sectors. It has also been argued that demand factors can lead to changing relative prices of the non-tradable goods (Froot and Rogoff, (1991), De Gregorio, Giovannini, and Wolf (1994)). However, previous empirical evidence showed that even within the category of tradable goods, the law of one price can fail across countries (Isard (1977), Giovannini (1988)). In particular, Engel (1993) and Rogers and Jenkins (1995) found that the short-run variance of real exchange rate is mainly explained by changes in the relative price for traded goods.

Recent studies have focused on the natures of such failures of the law of one price. One well accepted explanation is that of short-term price rigidities, due, say, to menu costs in changing prices (Mussa (1986), Ghosh and Wolf (1994)). The large empirical studies on pricing-to-market may present another explanation (see, among others, Knetter (1989), Marston (1990), Ohno (1989)).

On the other hand, Engel and Rogers (1994) found that national borders are a significant contributor to cross-country price variability beyond transporting costs. They also showed that the variability in the price of goods in two different locations depends on the distance between locations, as in gravity models of trade (see also Wei and Parsley (1995)). In addition, Engel and Rogers (1995) found that the law of one price holds more nearly for country pairs that are within geographic regions than the country pairs that are not.

There are several reasons why there may be smaller price disparities intraregionally First, two locations within a region are usually closer together than interregionally. When goods are costly to transport between than two locations in different regions. locations, arbitrage may not fully equalize prices. Since two locations within a region are geographically close, shipping costs within a region are smaller. Second, two locations within a region frequently form free trade areas, customs unions or common markets and may share common distribution systems. The absence of trade barriers clearly could help to explain why the law of one price holds more nearly within regions. Third, the nominal exchange rate between locations within a region is usually less Thus, if the nominal prices are sticky in the currency of the country in variable. which the final goods is sold, price disparities are smaller intraregionally than interregionally.

The purpose of this paper is to examine the nature of the deviations from the law of

one price in an East Asian "region". In particular, we investigate how "prices" in East Asian economies (that is, Korea, Taiwan, Hong Kong, Singapore, Malaysia, Indonesia, Thailand, and the Philippines) were correlated with those in Japan and with those in the United States. The analysis of "prices" in East Asian economies is particularly noteworthy for several reasons.

First of all, East Asian economies are geographically closer to Japan than to the United States. Thus, if the spread of the cross-border prices is explained by the geographical distance between countries, the law of one price would hold more nearly between Japan and East Asian economies than between the United States and East Asian economies. Second, Japan as well as the United States are the biggest trade partners for East Asian economies, especially in their imports (see Table 1). In addition, the amount of direct investment from Japan to East Asian economies has drastically increased during the last decade (see Table 2). These factors can also make the law of one price hold more nearly between Japan and East Asian economies.

However, East Asian currencies have traditionally been more tied to the US dollar than to the Japanese yen. Among East Asian economies, Hong Kong have been pegging its currency to the US dollar since the mid 1980s. On the other hand, most of other East Asian currencies have adopted the basket system to peg their currencies. Although the composition of the basket is not officially announced, it is well known that East Asian currencies have been strongly tied to the US dollar and that the weight of the yen has not been significant (see Frankel (1991) and Frankel and Wei (1993)). In addition, the US dollar is still a dominant invoice currency in East Asian foreign trade (see, for example, Tavlas and Ozeki (1992), Ito (1993), Fukuda and Ji (1995), Kawai (1996)). If the price of a consumer good might be sticky in local currency, these factors probably imply that the cross-border prices would fluctuate large between East Asia and Japan, but would be stable between East Asia and the United States.

In this paper, we analyze two different types of "prices". One is overall price levels in terms of the same currency. Using both consumer price and wholesale price indexes, we investigate how the overall price levels are correlated across countries. Needless to say, the correlation needs to be one if PPP holds exactly. The other is relative prices between different commodities within a country. Constructing the relative prices through dividing each commodity's price index by the overall price index, we explore how the relative prices move together across countries. Since the nominal exchange rate does not appear in the construction, the relative prices are not directly related to the issues on PPP. However, the cross-country correlation of the relative prices can be another good indicator to see how prices are linked internationally.

In our empirical results, we demonstrate that the overall price levels in East Asian economies were closely related to those in the United States. However, we also show that the relative prices in East Asian economies, especially in Taiwan and Korea, were closely correlated to those in Japan. The results are in marked contrast with the price correlation in other regions. For example, in north America, prices in Canada have had strong linkage with those in the United States both in the overall price levels and in the relative prices. A similar result will probably hold true among European countries. Thus, our results imply that East Asian economies have had a special international price linkage with Japan and the United States.

There are several previous studies which investigated price linkage in East Asia. Among others, Goto and Hamada (1994) applied principal component analysis to evaluate the degree of confluence of consumer price changes within East Asia. Similarly, Taguchi (1994) and Kawai and Okumura (1996) examined intraregional correlation of domestic inflation rates in East Asia. However, they focused neither on the issues on PPP nor on the international linkage of the domestic relative prices.

The paper proceeds as follows. The next section shows that the real exchange rates of East Asian economies were more stable to the United States than to Japan during the post-Bretton-Woods period. Section 3 extends this result by examining the stability of the East Asian real exchange rates after 1987 and during the Bretton-Woods period. Section 4 shows that the relative prices in East Asian economies were more closely correlated with those in Japan rather than with those in the United States. Section 5 tentatively considers why the relative prices, but not the overall price levels, have had a strong linkage between East Asian economies and Japan. Section 6 summarizes our main results and refers to a possible extension.

2. Real Exchange Rates in East Asia

This section investigates the behavior of real exchange rates for eight East Asian economies (that is, Korea, Taiwan, Hong Kong, Singapore, Malaysia, Indonesia, Thailand, and the Philippines). Specifically, we examine whether real exchange rates of these East Asian economies are more stable to Japan or to the United States. The series in calculating real exchange rates consist of the overall consumer price and overall wholesale price indexes. The price data are monthly, beginning in April 1973 and running to March 1996 in CPI and to December 1995 in WPI.¹ These

The WPI data in Hong Kong is not available. The WPI in Singapore starts from January 1974 and the WPI data in Malaysia starts from January 1986.

periods correspond to the time in which the yen/dollar nominal exchange rate was floating.

Before discussing the stability of real exchange rates, we first examined whether real exchange rates are stationary in East Asian economies. As is well known, individual tests on nominal exchange rates and price levels usually reveal the presence of a unit root, i.e. that they are integrated of order 1 [I(1)]. However, if PPP holds exactly, the real exchange rate is a stationary series with a mean equal to zero. In addition, given that price levels in terms of the same currency unit are I(1), a necessary condition for PPP is that these price levels are cointegrated.²

We first applied the unit root tests to real exchange rates of East Asian economies in terms of the US dollar and in terms of the Japanese yen. For the log of the real exchange rates, we performed three different types of unit root tests: the Augmented Dickey-Fuller (tau) by Dickey and Fuller (1979), Phillips-Perron (Z) by Phillips and Perron (1988), and Weighted Symmetric (Pantula et al. (1994)). In any type of test, test statistics was computed both with and without time trend. In addition, we selected the number of lags so as to maximize AIC2 in each test.

Table 3 reports the results of our unit root tests. Regardless of the choice of price index or time trend, almost all of unit root tests could not reject the null hypothesis that the real exchange rates have a unit root for eight East Asian economies. in the case of CPI, we could not reject the null hypothesis at 10% significance level except for one case (i.e., the Weighted Symmetric test with trend for Taiwan's real In the case of WPI, the results were exchange rates in terms of the Japanese yen). For example, we could reject the null hypothesis at 5% significance slightly modest. level for all unit root tests in Philippine real exchange rates in terms of the Japanese yen without trend and in Philippine real exchange rates in terms of the US dollar with trend. However, even in the case of WPI, almost all of other tests could not reject the null hypothesis that the real exchange rates have a unit root at reasonable significance level. Thus, we can generally conclude that PPP does not hold in East Asian economies in the sense that real exchange rates are not stationary both in terms of the US dollar and in terms of the Japanese yen.

However, even if (the log of) the real exchange rate has a unit root, the degrees of deviations from PPP can be different depending on whether we measure it in terms of the US dollar or in terms of the Japanese yen. To see this, we calculated both the

² For a broad survey on previous studies which investigated the stationarity of real exchangerates, see section 2 in Froot and Rogoff (1994).

unconditional and conditional variances of real exchange rates in terms of both the US dollar and the Japanese yen for eight East Asian economies. In calculating the variances of the logged real exchange rates, we used their first-differences because they have a unit root.

We examined the unconditional variance of a series because it may capture the long-term volatility of real exchange rates. However, in order to capture the short- to medium-term volatility of real exchange rates, we also measured the variance of the news about the series. Specifically, we estimated sixth-order autoregressions for the first-difference of each logged real exchange rate and then calculated the volatility measure as the variance of its forecast error.

Table 4 reports the calculated variances of the real exchange rates. For all CPI's and WPI's, the variances in terms of the US dollar were smaller than those in terms of the Japanese yen both in the unconditional and conditional variances. For most cases, the variances in terms of the Japanese yen were more than double of those in terms of the US dollar. In particular, for CPI's in Korea, Hong Kong, and Singapore, the real exchange rates in terms of the Japanese yen fluctuated more than four times as large as those in terms of the US dollar. Thus, we can see that for all East Asian economies, the degrees of deviations from PPP were much smaller in terms of the US dollar than in terms of the Japanese yen.

The above results imply that in East Asian economies, the geographical distance between countries is not a key factor in explaining the spread of the cross-border price levels. Instead, noting that East Asian currencies have been strongly tied to the US dollar, sticky prices of final goods in local currency can explain well the degrees of deviations from PPP. It is noteworthy that the result similarly holds well even in a small open economy such as Hong Kong and Singapore. This indicates that even if tariff rates and non-tariff barriers are negligible, there is little evidence of price arbitrage in final goods.³

3. Real Exchange Rates during Alternative Periods

Lee and Swagel (1994) reported tariff rates and non-tariff barriers in various countries. According to their report, both tariff rates and non-tariff barriers are zero in Hong Kong, tariff rates range from 0.6 to 0.9 and non-tariff barriers 1.1 to 1.2 in Singapore. These figures are much lower than those of OECD countries. For example, they reported that tariff rates range from 4.5 to 4.9 and non-tariff barriers from 26.6 to 37.5 in the United States.

In the last section, we showed that real exchange rates of East Asian economies were more stable to the United States than to Japan during the post-Bretton-Woods period when the yen/dollar nominal exchange rate was floating. This section investigates the validity of the results by extending the analysis to two directions.

First, we examined the stability of the East Asian real exchange rates after 1987. We picked up this sample period because during this period, the economic linkage between East Asia and Japan has increased a lot, especially in direct investment from Japan (see Table 3). For this short sample period, the unit root tests may not be powerful in general. However, almost all of our unit root tests could not reject the null hypothesis that the logged real exchange rates have a unit root in eight East Asian economies. Thus, taking their first-differences, we calculated both the unconditional and conditional variances of real exchange rates in terms of both the US dollar and the Japanese yen.⁴

Table 5 reports the computed variances of the real exchange rates. Both in the unconditional and conditional variances, the variances in terms of the Japanese yen were larger than those in terms of the US dollar in all East Asian economies. In particular, we can see that the variances in terms of the Japanese yen are almost the same as those in Table 4 but that the variances in terms of the US dollar are smaller than those in Table

4. This indicates that after the mid 1980s, the degrees of deviation from PPP became relatively larger in terms of the Japanese yen rather than in terms of the US dollar for all East Asian economies.

Our second extension is to examine the stability of the East Asian real exchange rates during the Bretton-Woods period. This period corresponds to the time in which the yen/dollar nominal exchange rate was fixed.⁵ If the nominal exchange volatility is a key contributor to deviations from PPP, we would find no significant difference in the stability between East Asian real exchange rates in terms of the US dollar and those in terms of the Japanese yen. Unless specified, the sample period is from January 1957 through July 1971. Depending on the availability of the data, we examined the relative stability of the real exchange rates for six East Asian economies (that is, Korea, Taiwan, Singapore, Malaysia, Thailand, and the Philippines).

We first applied the unit root tests to the logged real exchange rates of East Asian

⁴ The method of calculating the conditional variances is the same as that in the last section.

⁵ During this period, the nominal exchange rates in East Asian economies were also fixed to the US dollar except for infrequent changes in the official parity.

economies.⁶ Table 6 reports the results of our unit root tests. In the case of WPI, any type of unit root tests could not reject the null hypothesis that real exchange rates in terms of the US dollar and the Japanese yen have a unit root in eight East Asian economies. Similarly, in the case of CPI, all unit root tests could not reject the null hypothesis except for two cases (i.e. Phillips- Perron tests for Malaysian rates in terms of the yen with trend and for Singapore's rates in terms of the yen with trend). Thus, even during the Bretton-Woods period, we can generally conclude that PPP does not exactly hold in East Asian economies in the sense that real exchange rates are not stationary both in terms of the US dollar and in terms of the Japanese yen.

We next calculated both the unconditional and conditional variances of real exchange rates in terms of both the US dollar and the Japanese yen for eight East Asian economies. In calculating the variances of the logged real exchange rates, we used their first-differences because the logged real exchange rates have a unit root. In calculating their conditional variances, we estimated sixth-order autoregressions for the first-difference of each logged real exchange rate.

Table 7 reports the calculated variances of the real exchange rates. When we used CPI, the variances of real exchange rates in terms of the US dollar were still smaller than those in terms of the Japanese yen in Malaysia and Singapore. However, in other cases of CPI and in all cases of WPI, the variances of the real exchange rates in terms of the US dollar were as large as those in terms of the Japanese yen. Thus, we can conclude that during the Bretton-Woods period, the degrees of deviations from PPP in terms of the US dollar were as big as those in terms of the Japanese yen for almost all East Asian economies.

4. International Linkage of Relative Prices

In the last two sections, we showed that when the yen/dollar nominal exchange rate was floating, the real exchange rates of East Asian economies were more stable to the United States rather than to Japan. We also showed that when the yen/dollar nominal exchange rate was fixed, the real exchange rates in East Asian economies were almost equally stable to the United States and to Japan. These results imply that the volatility of nominal exchange rates was a significant contributor to cross-country price variability in East Asian economies during the post-Bretton-Woods period. However, these results do not imply that besides the effects of nominal exchange rate volatility,

In any type of test, test statistics was computed both with and without time trend. In addition, we selected the number of lags so as to maximize AIC2.

East Asian economies have had stronger "price" linkage to the US economy rather than to the Japanese economy.

The purpose of this section is to analyze whether relative prices among different commodities in East Asian economies were more similar to those in the United States than to those in Japan. The analysis is particularly noteworthy because we do not use nominal exchange rates to calculate relative prices within a country. If the price of a consumer good is simply sticky in local currency, the analysis would not find any special relationship among the relative prices of different countries. However, if there is some "price" linkage across borders, the analysis may find some correlations of the relative prices.

In the analysis, we calculated the relative prices among different commodities for eight East Asian economies, the United States, and Japan. We then examined whether the relative prices in the East Asian economies are more closely correlated with those in the United States or with those in Japan. We constructed each relative price through dividing each commodity's price index by the overall price index within a country. Specifically, for each commodity j and country k, we defined the relative price by

$$q_i^k = \log(p_i^k/p^k),$$

where p_i^k is the commodity i's price index in country k and p^k is the overall price index in country k.

We regressed the commodity i's relative price in each East Asian economy k, q_i^k , on the corresponding relative prices in Japan and the United States, q_i^{JAPAN} and q_i^{USA} as follows:

$$q_i^k = \text{constant} + a^* q_i^{JAPAN} + b^* q_i^{USA},$$

In the regression, we included a time trend, if any, and used the ML method if disturbances display the first-order serial correlation.⁷ The data are annual price indexes for the overall consumer price and wholesale price indexes and their subcategories. Unless specified, the sample period is from 1970 to 1994.

Table 8 reports the results of our regression. One striking result is that the relative prices in East Asian economies tended to be more closely correlated with those in Japan rather than with those in the United States. The tendency was most

We also included a squared time trend, if any.

conspicuous in Taiwan and Korea for which a number of subcategories are available. That is, for the Taiwan's relative prices, the coefficient of the Japan's relative price was much greater than that of the US relative price in seven of nine WPI commodities and in all of three CPI commodities. For the Korea's relative prices, the coefficient of the Japan's relative price was significantly greater than that of the US relative price in ten of thirteen WPI commodities and in three of four CPI commodities.

The tendency was essentially the same in the Philippine price indexes and the Thailand's WPI, although the estimated coefficients were less statistically significant. That is, for the Philippine relative prices, the coefficient of the Japan's relative price was greater than that of the US relative price in all of two WPI commodities and in two of three CPI commodities. For Thailand's WPI, the coefficient of the Japan's relative price was significantly greater than that of the US relative price in four of six commodities.

The tendency was less clear in Hong Kong, Singapore, Indonesia, Malaysia, and Thailand's CPI. In particular, except for Thailand's CPI, the relative prices of food in CPI were more correlated with those in the United States rather than those in Japan. However, even in these economies, our overall results show that the coefficients of the US relative prices never dominate those of the Japan's relative prices. In addition, most of the estimated results were not statistically significant. We conjecture that the less significant estimations may have occurred because price indexes have large measurement errors in these economies. In fact, for these economies, the numbers of sub-categories we could obtain were very limited.

5. Interpretations

In previous sections, we showed that the overall price levels in East Asian economies were closely related to those in the United States. However, we also demonstrated that the relative prices in East Asian economies, especially in Taiwan and Korea, were closely correlated with those in Japan. These results look in marked contrast with price correlation in other regions. For example, in north America, prices in Canada have had strong linkage with those in the United States not only in the overall price levels but also in the relative prices.

Table 9 shows this for six selected commodities (that is, food, textile, wood, paper, non-metal, and basic metal). As in the last section, we regressed each commodity's relative price in Canada on the corresponding commodity's relative prices

in Japan and the United States.⁸ The data are monthly price indexes for the overall wholesale prices and their sub-categories from January 1975 to October 1994. Not surprisingly, the results in the table show that the relative prices in Canada were more closely correlated with those in the United States rather than with those in Japan. Since the overall price levels in Canada have also strong linkage with those in the United States, this implies that prices in Canada have had strong linkage with those in the United States both in the overall price levels and in the relative prices.

The result, however, leads to a question why the relative prices, but not the overall price levels, have had a strong economic linkage between East Asia and Japan. The first possible answer is that East Asian economies and Japan have a similar economic structure, including similar endowments and preferences. In this case, their autarky relative price structure can be similar. Thus, even without any overall price arbitrage, the relative prices in East Asia can have had a strong correlation with those in Japan.

The second possible answer is that prices adjust to the nominal exchange rate changes with a very low speed but the adjustment speeds are almost the same for most of commodities. In this case, the overall price levels are not equalized across the national borders for a long time. However, the relative prices can move together across the borders.

The third possible answer is that there exist little price arbitrage in final goods but large price linkage in intermediate goods between Japan and East Asian economies. Without price arbitrage in final goods, it is natural that the overall price levels are not equalized across the national borders. However, if prices of intermediate goods are correlated, it is possible that the relative prices move together across the borders.

6. Concluding Remarks

Not a few previous studies concluded that the law of one price holds more nearly for country pairs that are within geographic regions than the country pairs that are not because of transporting costs. The results of our paper showed that this conclusion does not carry through to the overall price levels in East Asian economies. That is, real exchange rates of East Asian economies were more stable to the United States rather than to Japan during the post-Bretton-Woods period. However, our results also showed that East Asian relative prices among different commodities were more

We also included a constant term and a time trend, if any, and used the ML method if disturbances display serial correlation.

correlated with those in Japan rather than with those in the United States. The results are inconsistent with a simple story of price rigidities due to menu costs because the price rigidities hardly explain international correlation of relative prices among different commodities. They are also inconsistent with the Balassa-Samuelson effect because the deviations from PPP were large among countries where their relative prices within countries moved similarly.

Of course, our results say little about why the relative prices, but not the overall price levels, have had a strong correlation between East Asia and Japan. Theoretically, various types of answers are possible. Under one extreme situation, the relative prices may have a strong correlation because the autarky relative price structures are similar between East Asia and Japan. In this case, our results do not necessarily imply any causal international price linkage between East Asia and Japan. However, without direct price linkage, similar autarky relative price structures tend to be formed through various indirect economic linkage among countries. Thus, even if the correlation comes from the similarity in their autarky relative price structures, it is consistent with the indirect price linkage between East Asia and Japan.

For example, there exist various another possible answers. In addition, suppose that price arbitrage in final goods is small but price linkage in intermediate goods is large. Without price arbitrage in final goods, it is natural that the overall if prices of price levels are not equalized across national borders. However, intermediate goods are correlated, it is possible that the relative prices move together Theoretically, cross-border price arbitrage is more likely to be across borders. undertaken at the intermediate goods level than at the consumer goods level. Thus, the closer together two regions are, the less dispersion is likely in the intermediate Needless to say, this possible explanation needs formal empirical goods prices. studies by using detailed data.

Data Appendix

Except for sub-categories of price indexes, all data of price indexes and exchange rates are taken from <u>International Financial Statistics</u>, IMF, various issues, and <u>Financial Statistics</u>, Taiwan <u>District</u>, The Republic of China, The Central Bank of China, various issues.

Sources of sub-categories of price indexes for each country are as follows:

- (1) USA: <u>Statistical Abstract of the United States</u>, US Department of Commerce, Economics and Statistics Administration Bureau of the Census, various issues.
- (2) Japan: <u>Annual Report on the Consumer Price Index</u>, Statistics Bureau, Management and Coordination Agency, Government of Japan. <u>Price Indexes Annual</u>, Research and Statistics Department, Bank of Japan.
- (3) Korea: <u>Price Statistics Summary</u>, Bank of Korea. <u>Economic Statistics Yearbook of Korea</u>, Economic Research Institute of Korea.
- (4) Taiwan: <u>Statistical Yearbook of the Republic of China</u>, Directorate-General of Budget, Accounting and Statistics, Executive Yuan, Republic of China.
- (5) Singapore: Yearbook of Statistics Singapore, Department of Statistics, Singapore.
- (6) Malaysia: Yearbook of Statistics Malaysia, Department of Statistics, Malaysia.
- (7) Indonesia: <u>Statistical Yearbook of Indonesia</u>, BIRO PUSAT STATISTIK JAKARTA. <u>Statistical Pocketbook of Indonesia</u>, BIRO PUSAT STATISTIK JAKARTA.
- (8) Thailand: <u>Statistical Yearbook Thailand</u>, National Statistical Office, Office of the Prime Minister.
- (9) The Philippine: <u>Philippine Statistical Yearbook</u>, National Statistical Coordination Board, Philippine.
- (10) Hong Kong and the others: <u>Statistical Yearbook for Asia and The Pacific</u>, the United Nations.

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Table 1 The Degrees of Trade Dependency on the USA and on Japan

1981 1986 1991 1994		- 1	בן בן
26. 2 37. 2 24. 5 21. 6	Export	NIFs Total	The II S (%)
16.9 16.1 16.4 14.3	Export Import Export	otal	9
26. 8 40. 1 4 25. 3 21. 3	Expor	Korea	
23 20 23 23	t Impor		
. 2 36.1 . 7 47.7 . 2 29.3 . 9 26.4	Import Export	Taiwan	
22. 22. 22. 20.	t Import	n	
5 13 5 23 5 19 1 18	t Expor	Singapore	
	t Impor	pore	
27. 31. 22. 20.	t Export Import Export	Hong Kong	
3 8. 4 7 7. 6 8 6. 4	I,	l	
4 17.7 4 20.1 6 18.4 4 22.2	E.	ASEAN T	
1 17.0 1 13.8 2 14.0	t Impor	Total	
18. 21. 23.	Expor	Thailand	
1 14.3 1 10.5 2 10.5 3 11.4	15	and	
35. 38.	t Export	Phili	
6 24.8 9 20.8 6 18.5	a cc o	Philippines	
8 19.6 8 12.0 5 16.8	ני	Indonesia	-
6 7.2 0 13.2 8 10.1	3 13 5	les1a	
	13 5 13 1	Malaysia	W-1 2
	1 14.6	+ 1 mnor	2:5
<u>σω</u> ∞	တြင	+	

1981 1986 1991 1994		For J
981 10. 50 23. 30 16. 50 24. 40 10. 90 28. 00 10. 10 18. 80 4. 70 23. 10 33. 40 20. 00 13. 10 10. 10 18. 80 4. 70 20. 40 28. 30 24. 10 14. 20 14. 20 14. 20 26. 10 17. 20 26. 10 17. 20 25. 90 12. 10 30. 00 8. 70 21. 40 5. 40 16. 40 22. 90 26. 10 17. 20 29. 40 27. 80 18. 60 27. 80 18. 60 27. 80 18. 60 18. 60 27. 80 18. 60 18. 60 27. 80 18. 60 18. 60 27. 80 18. 60 18. 60 18. 60 27. 80 18. 60 <t< td=""><td>Export Import Export Import Export Import Export Import</td><td>Japan(%)</td></t<>	Export Import Export Import Export Import Export Import	Japan(%)
23. 30 26. 90 22. 70 20. 90	[mport]	1
16. 50 15. 60 17. 20 14. 10	Export	Korea
24. 40 34. 40 25. 90 24. 80	Import E	
10. 90 11. 40 12. 10 11. 10	xport	Taiwan
28. 00 34. 10 30. 00 27. 50	mport E	τς.
10. 10 8. 60 8. 70 7. 00	xport I	Singapore
18. 80 19. 90 21. 40 21. 90	mport E	
4. 70 4. 70 5. 40 5. 00	xport I	Hong Kong
23. 10 20. 40 16. 40 14. 00	mport E	
28. 30 2 22. 90 2 18. 60 2	t Export Import	ASEAN Total
24. 10 1 26. 10 1 27. 80 1	port Ex	1 1
4. 20 2 7. 90 2 8. 00 3	port Im	Thailand
6. 40 1 9. 10 2 0. 60 1	port Ex	Ph
7. 70 1' 0. 20 20 5. 00 2	port 1m	Philippines
7. 00 44 0. 00 36 4. 20 3	port Exi	es Inc
1. 90 15 3. 90 24 3. 90 27	. 90 30	Indonesia
14. 20 26. 40 17. 70 17. 00 44. 90 15. 10 23. 30 20. 50 17. 90 29. 10 20. 20 20. 00 36. 90 24. 50 15. 90 26. 10 18. 00 30. 60 15. 00 24. 20 30. 90 27. 60 11. 90 26. 70	t Export Import Import Export Import Import Export Import	Mal
. 30 20. . 90 26. . 90 26.	10 24.	Malaysia
50 10 70	40	1

Sources) IMF, Direction of Trade Statistics, various issues.

Table 2 Sources of Inward Direct Investment in East Asian Economies

1/.	6.6	4.4	46.8	<u>51</u> .3	21.1	24.0	32.5	Japan	
13.7	4.1	28.8	13.5	22.6			23.6	The U.S.	1994
6.02	10.6		23.2	44.0			16.2	Japan	
	3.1		18.5	16.1	39.4	33.0	21.2	The U.S.	1991
	49.9		20.5	19.6		32.9	38.9	Japan	
2.0	13.3		19.	22.5		18.0	35.4	The U.S.	1986
3 \{\bar{2} \\ 3 \}	13 N		8.02	31.5			23.8	Japan	
Z	N/A	46.9	11.6	43.6		N/A	58.6	The U.S.	1981
Malaysia	Indonesia	Philippines Indonesia	haila	Hong	Singapore	Taiwan	Korea		

Notes) 1. Figures in the table denote that the weights of the USA and Japan in the total inward direct investment to East Asian Economies.
2. The data in Hong Kong is based on surveys on manufafturing industry.
Sources) Jetro, White Paper, various issues.

Table 3 Unit Root Tests During the Post-Bretton-Woods Period

(1) CPI real exchange rates with no trend

country country	nange rae	in the US	dollar		in the Ja		
sample period		TestStat		lags	TestStat	P-value	lags
Korea	WS	-1. 472	0. 532	3	-1.739	0. 338	3
73:4-96:3	ADF	-1. 611	0.478	3	-1.565	0.501	5
10.1 00.0	PP	-4. 033	0. 535	3	-4. 621	0.471	5
Hong Kong	WS	-0. 563	0. 945	4	B	0. 545	5
73:4-96:3	ADF	-0. 208	0. 938	4	-2. 030	0. 273	5
	PP	-0.093	0. 952	4		0. 503	5
Indonesia	WS	-1. 173	0. 736	3	1	0. 755	13
73:4-96:3	ADF	-0. 645	0.860	3		0.762	13
1011 0010	PP	-0. 933	0.892	3	-0.914	0. 894	13
Malaysia	WS	-1. 146	0.751	4	-0. 792	0. 897	7
73:4-96:3	ADF	-1. 368	0. 597	7	-1.285	0. 636	7
10.1 00.0	PP	-1. 803	0.803	7	-1.494		7
the Philippines	WS	-1. 906	0. 237	11		0. 385	
73:4-96:3	ADF	-1. 633	0.466	11	-1. 535		
	PP	-6. 054	0.343	11	-3. 867	Name and Address of the Owner, where the Person of the Owner, where the Person of the Owner, where the Person of the Owner, where the Owner, which is the Owner, which	
Thailand	WS	-1. 830	0. 281	3	-1. 144		
73:4-96:3	ADF	-1. 505	0. 531	3	-1. 169		
70.1	PP	-3. 343	0.616	3	-1.976		
Singapore	WS	-1. 406	0. 580	7			
73:4-96:3	ADF	-1. 638	0.464	3	-1. 997		
	PP	-3. 464	0.601	Ç	-3.901		
Taiwan	WS	-1. 949	0. 214	13			
73:4-96:3	ADF	-1.997	0. 288	13	-2. 139		
	PP	-3. 901	0. 550	13	3 -7. 559	0. 243	3

(2) CPI real exchange rates with trend

(2) CPI real exc	nange rat						
country		in the US				panese ye	
sample period		TestStat P	-value	lags	TestStat		lags
Korea	WS	-1. 691	0.823	3	-2. 699	0. 186	3
73:4-96:3	ADF	-1.625	0.782	3	-2. 770	0. 208	5
, , , ,	PP	-4.170	0.877	3	-12.369	0. 297	5
Hong Kong	WS	-0. 450	0. 995	4	-2.230	0. 481	5
73:4-96:3	ADF	0. 226	0.996	4	-2.030	0. 585	5
0.1 00.0	PP	0. 376	0. 998	4	-7. 936	0. 589	5
Indonesia	WS	-0. 727	0. 988	3	-1. 626	0. 849	13
73:4-96:3	ADF	-2. 566	0. 296	3	-2. 679	0. 245	13
10.1 00.0	PP	-10. 188	0.424	3	-12. 158	0. 308	13
Malaysia	WS	-1.853	0.741	4	-2. 076	0. 595	7
73:4-96:3	ADF	-1. 296	0.889	7	-2. 024	0. 589	7
10.1 00.0	PP	-9. 630	0.461	7	-8. 897	0. 515	7
the Philippines	WS	-1.554	0.875	11	-2. 364	0. 383	3
73:4-96:3	ADF	-1. 188	0.913	11	-2. 437	0.360	3
70.1 00.0	PP	-6. 912	0.672	11	-11.618	0. 337	3
Thailand	WS	-1. 248	0. 946	5	-2. 163	0. 531	3
73:4-96:3	ADF	-1.857	0.677	3	-2. 313	0.427	5
70.1 00.0	PP	-7. 079	0.658	3	-9.147	0. 496	5
Singapore	WS	-1. 153	0. 959	7	-2. 388	0. 367	5
73:4-96:3	ADF	-1. 006	0.943	3	-2. 221	0.478	5
	PP	-2. 915	0. 940	3	-7. 246	0.644	
Taiwan	WS	-1. 921	0. 700	13	-2. 664	0. 202	
73:4-96:3	ADF	-1. 848	0. 681		-3. 044	0. 120	
10.1 50.0	PP	-3. 959	0. 889		1	0. 248	3

Table 3 Unit Root Tests During the Post-Bretton-Woods Period (continued)

(3) WPI real exchange rates with no trend

country		in the US	dollar			apanese y	
sample period		TestStat	P-value	lags	TestStat		lags
Korea	WS	-1. 612	0. 428	12	-1. 652		3
73:4-95:12	ADF	-2. 865	0.050	2	-1. 375		3
	PP	-8. 223	0. 208	2	-3. 513		3
Indonesia	WS	-1.406	0. 581	4			5
73:4-95:12	ADF	-1. 448	0. 559	2	-0. 923		7
	PP	-3.763	0. 566	2	-1. 595		7
the Philippines	WS	-1. 151	0.748	10	-2. 819		3
73:4-95:12	ADF	-1. 632	0.467	10	-2. 903		3
	PP	-14.347	0.047	10	-14.227	_	3
Thailand	WS	-1.904	0. 238	5	-1. 502		3
73:4-95:12	ADF	-1. 801	0.380	3	-1.320		3
	PP	-4. 826	0. 451	3	-3. 007		3
Taiwan	WS	-2. 058	0. 165	10	-1. 963		
73:4-95:12	ADF	-2.376	0. 149	6	-1. 746		3
	PP	-12. 079	0.082	6	-5. 694	0.372	
Singapore	WS	-1.107	0.772	9	-1.335		
74:1-95:12	ADF	-1. 650	0. 457	9	-1.392		
	PP	-6. 389	0.318	9	-3. 029	0. 654	
Malaysia	WS	-1.407	0. 580	4	-1. 160		
86:1-95:12	ADF	-1.290	0.634	3	-3. 279	0. 016	
	PP	-4. 32687	0.50214	3	-23. 167	0.005	3

(4) WPI real exchange rates with trend

country		in the US	dollar		in the Ja		en
sample period		TestStat		lags	TestStat	P-value	lags
Korea	WS	-1. 524	0. 884	12	-1. 956	0. 678	3
73:4-95:12	ADF	-3. 029	0. 124	2	-2. 989	0. 135	3
	PP	-8. 553	0. 541	2	-10.504	0.403	3
Indonesia	WS	-0.782	0. 986	4	-1. 295	0. 938	5
73:4-95:12	ADF	-3. 745	0.020	2	-2.515	0. 321	7
	PP	-14.538	0. 202	2	-9. 909	0. 442	7
the Philippines	WS	-4.578	0.001	10	-2. 799	0. 145	3
73:4-95:12	ADF	-4. 496	0.002	10	-3. 342	0.060	3
	PP	-35. 346	0.003	10	-17.717	0. 111	3
Thailand	WS	-1. 682	0. 827	5	i	0. 436	3
73:4-95:12	ADF	-1. 560	0.808	3		0. 331	3
	PP	-5. 170	0. 809	3		0. 315	3
Taiwan	WS	-2. 269	0. 452	10	1	0. 140	5
73:4-95:12	ADF	-2. 701	0. 236	6		0. 082	3
	PP	-12. 952	0. 269	6		0. 174	3
Singapore	WS	-2. 349	0. 394	3	4	0. 197	5
74:1-95:12	ADF	-2. 419	0. 370	3		0. 291	5
	PP	-10.454	0. 406	3			
Malaysia	WS	-1. 697	0. 820			0. 399	
86:1-95:12	ADF	-1. 693	0. 754	4			
	PP	-9. 927	0. 441	4	-27.953	0. 014	3

Notes) 1. "WS", "ADF", and "PP" respectively denote unit root tests of Weighted Symmetric, Augumented Dickey-Fuller (tau), and Phillips-Perron (Z).

2. P-values of ADF and PP were calculated by MacKinnon (1994). P-values of WS were based on TSP v. 4.3.

(1) CPI real exchange rates relative values (B) variances (A) variances type of country (B)/(A)in the yen sample period in US dollar variances 4. 269 0.00098 0.00023 unconditional Korea 4. 212 0.00088 0.00021 conditional 73:4-96:3 4. 354 0.00091 0.00021 unconditional Hong Kong 4.210 0.00079 0.00019 conditional73:4-96:3 1. 596 0.00158 0.00099 unconditional Indonesia 1.556 0.00142 0.00092 conditional 73:4-96:3 3. 412 0.00076 0.00022 unconditional Malaysia 3. 133 0.00063 0.00020 conditional 73:4-96:3 3.008 0.00166 0.00055 unconditional the Philippines 2.877 0.00155 0.00054 conditional 73:4-96:3 4. 543 0.00082 0.00018 unconditional Thailand 4.420 0.00075 0.00017 conditional 73:4-96:3 3.086 0.00070 0.00023 Singapore unconditional 3.209 0.00063 0.00019 <u>conditional</u> 73:4-96:3 2. 864 0.00099 0.00035 unconditional Taiwan 2.832 0.000890.00031 conditional 73:4-96:3

(2) WPI real ex	change rates			
country	type of	(A) variances	(B) variances	relative values
sample period	variances	in US dollar	in the yen	(B) / (A)
Korea	unconditional	0. 00027	0. 00079	
73:4-95:12	conditional	0. 00025	0. 00074	2. 933
Indonesia	unconditional	0. 00080	0. 00127	1, 586
73:4-95:12	conditional	0. 00077	0. 00123	
Malaysia	unconditional	0. 00017	0. 00062	
86:8-95:12	conditional	0. 00016	0. 00055	
the Philippines	unconditional	0. 00053	0. 00123	I
86:8-95:12	conditional	0. 00051	0. 00115	
Thailand	unconditional	0. 00025	0. 00066	
86:8-95:12	conditional	0. 00024	0. 00062	
Singapore	unconditional	0. 00030	0. 00061	2. 039
74:8-95:12	conditional	0. 00026	0. 00051	1. 993
Taiwan	unconditional	0. 00022	0. 00065	1
73:4-95:12	conditional	0. 00019	0. 00059	3. 041

Table 5 The Variances of the Real Exchange Rates after 1987

(1) CPI real exchange rates

country sample period	type of variances	(A) variances in US dollar	(B) variances in the yen	relative values (B)/(A)
Korea	unconditional	0, 00008	0. 00084	10. 745
87:1-96:3	conditional	0. 00006	0.00071	12. 175
Hong Kong	unconditional	0. 00002	0. 00090	38. 063
87:1-96:3	conditional	0. 00002	0. 00073	35. 764
Indonesia	unconditional	0. 00005	0. 00082	1
87:1-96:3	conditional	0. 00004	0. 00069	
Malaysia	unconditional	0. 00013	0. 00081	6. 378
87:1-96:3	conditional	0. 00011	0. 00068	
the Philippines	unconditional	0. 00025	0. 00136	1
87:1-96:3	conditional	0. 00021	0. 00106	
Thailand	unconditional	0. 00005	0. 00064	
87:1-96:3	conditional	0. 00004	0. 00054	
Singapore	unconditional	0. 00011	0. 00063	· I
87:1-96:3	conditional	0. 00011	0. 00055	
Taiwan	unconditional	0. 00026		3. 292
87:1-96:3	conditional	0. 00023	0. 00071	3. 024

(2) WPI real exchange rates

(2) WP1 real ex		(4)	(D) variance	relative values
country	type of	(A) variances	(B) variances	V.
sample period	variances	in US dollar	in the yen	(B) / (A)
Korea	unconditional	0. 00007	0. 00063	8. 989
87:1-95:12	conditional	0. 00006	0. 00054	9. 477
Indonesia	unconditional	0. 00012	0. 00058	•
87:1-95:12	conditional	0. 00009	0. 00049	
Malaysia	unconditional	0. 00017	0. 00062	i
87:1-95:12	conditional	0. 00016	0. 00055	
Philippines	unconditional	0. 00037	0. 00116	
87:1-95:12	conditional	0. 00032	0. 00096	
Thailand	unconditional	0. 00010	0. 00050	
87:1-95:12	conditional	0. 00009	0. 00039	
Singapore	unconditional	0. 00039	0. 00063	
87:1-95:12	conditional	0. 00034	0. 00058	
Taiwan	unconditional	0. 00009	0. 00060	
87:1-95:12	conditional	0. 00008	0. 00050	6. 049

Table 6 Unit Root Tests during the Bretton-Woods Period

(1) CPI real exchange rates with no trend

country		in the US			in the Ja		en
sample period		TestStat	P-value	lags	TestStat	P-value	lags
Malaysia	WS	1. 696	1. 000	2	1. 913	1. 000	2
57:1-71:7	ADF	0. 320	0. 978	2	0. 772	0. 991	6
	PP	0. 503	0. 978	2	0. 385	0. 974	6
the Philippines	WS	-1. 150	0. 748	2	-0. 699	0. 920	2
57:1-71:7	ADF	-1. 381	0. 591	2	-1.012	0.749	2
0111 / 21/	PP	-3. 944	0. 545	2	-2. 069	0.772	2
Taiwan	WS	-1. 693	0. 370	2	-1. 657	0. 396	2
57:5-71:7	ADF	-1. 796	0. 383	2	-1. 454	0. 556	2
00	PP	-6. 613	0.302	2	-5. 094	0. 425	2
Singapore	WS	0.360	0. 996	3	2. 285	1. 000	3
61:1-71:7	ADF	0. 829	0. 992	3	-0. 497	0. 893	5
	PP	0. 799	0. 986	3	-0. 505	0. 926	5

(2) CPI real exchange rates with trend

(2) CPI real exc	nange rau	tes with t	rena				
country		in the US	dollar		in the Ja	ipanese y	en
sample period		TestStat	P-value	lags	TestStat	P-value	lags
Malaysia	WS	-1.460	0. 902	2	-2. 620	0. 224	2
57:1-71:7	ADF	-1. 204	0. 910	2	-3. 040	0. 121	2
	PP	-3. 881	0.893	2	-20. 917	0. 059	2
the Philippines	WS	-2. 762	0. 159	2	-2. 580	0. 246	2
57:1-71:7	ADF	-2. 563	0. 297	2	-2. 428	0. 365	2
	PP	-12. 617	0. 285	2	-10.873	0. 380	2
Taiwan	WS	-2. 453	0. 323	2	-2. 585	0. 243	2
57:5-71:7	ADF	-2. 225	0.476	2	-2. 543	0. 307	2
	PP	-11. 734	0. 330	2	-13.085	0. 263	2
Singapore	WS	-1. 143	0. 960	3	-3. 160	0. 054	2
61:1-71:7	ADF	-1. 130	0. 924	3	-2. 260	0.456	5
c c	PP	-4. 207	0.874	3	-23. 776	0. 033	5

Table 6 Unit Root Tests During the Bretton-Woods Period (continued)

(3) WPI real exchange rates with no trend

country		in the US			in the Ja		en
sample period		TestStat	P-value	lags	TestStat		lags
Korea	WS	-0. 960	0. 840	2	-1. 213	0.711	2
57:1-71:7	ADF	-1.810	0. 375	2	-1.749	0. 406	2
	PP	-5. 093	0. 425	2	-5. 100	0. 424	2
the Philippines	WS	-1. 429	0. 564	2	-1. 662	0. 392	2
57:1-71:7	ADF	-1. 444	0. 561	2	-1. 490	0. 538	2
	PP	-5. 280	0.408	2	-5. 637	0. 377	2
Thailand	WS	-2.580	0.040	2	-2. 297	0. 088	2
57:1-71:7	ADF	-2.394	0. 143	2	-2. 414	0. 138	3
0,72 ,21,	PP	-12. 763	0.069	2	-12. 361	0. 076	3
Taiwan	WS	-1.696	0. 368	2	-1. 204	0.717	2
57:5-71:7	ADF	-1. 602	0.483	2	-2. 065	0. 259	2
	PP	-4. 766	0. 457	2	-6. 421	0. 316	2

(4) WPI real exchange rates with trend

(4) WPI real exc	nange ra						
country		in the US	dollar		in the Ja		
sample period		TestStat	P-value	lags	TestStat	P-value	lags
Korea	WS	-2. 469	0. 312	2	-2. 427	0. 340	2
57:1-71:7	ADF	-2. 364	0.399	2	-2.225	0.476	2
	PP	-11. 581	0. 339	2	-10. 733	0.389	2
the Philippines	WS	-2. 828	0. 134	2	-2. 670	0. 200	2
57:1-71:7	ADF	-2. 684	0. 243	2	-2. 664	0. 251	2
	PP	-15. 408	0. 172	2	-14. 753	0. 194	2
Thailand	WS	-2.700	0. 186	2	-2. 807	0. 142	3
57:1-71:7	ADF	-2.507	0. 324	2	-2. 684	0. 243	3
	PP	-14. 637	0. 198	2	-16. 701	0. 135	3
Taiwan	WS	-1. 592	0. 862	2	-1. 916	0. 704	2
57:5-71:7	ADF	-1. 263	0. 897	2	-1. 868	0.671	2
e ·	PP	-4. 646	0.846	2	-8. 361	0. 555	2

Notes) 1. "WS", "ADF", and "PP" respectively denote unit root tests of Weighted Symmetric, Augumented Dickey-Fuller (tau), and Phillips-Perron (Z).

2. P-values of ADF and PP were calculated by MacKinnon (1994). P-values of WS were based on TSP v. 4.3.

(1) CPI real ex	change rates			
country	type of	(A) variances	(B) variances	relative values
sample period	variances	in US dollar	in the yen	(B)/(A)
Malaysia	unconditional	0. 00005	0. 00011	2. 244
57:8-71:7	conditional	0. 00005		
Philippines	unconditional	0. 00355	0. 00356	
57:8-71:7	conditional	0. 00350		
Taiwan	unconditional	0. 00042	0. 00050	
57:12-71:7	conditional	0. 00041	0. 00049	
Singapore	unconditional	0. 00012	0. 00017	
61:8-71:7	conditional	0. 00010	0. 00015	1. 441

(2) WPI real ex	change rates			
country	type of	(A) variances	(B) variances	relative values
sample period	variances	in US dollar	in the yen	(B)/(A)
Korea	unconditional	0. 00522	0. 00527	1. 010
57:8-71:7	conditional	0. 00515	0. 00520	
Phillipines	unconditional	0. 00377	0. 00378	1
57:8-71:7	conditional	0. 00371	0. 00372	
Thailand	unconditional	0. 00064	0. 00065	
57:8-71:7	conditional	0. 00061	0. 00063	
Taiwan	unconditional	0. 00022	0. 00022	
58:5-71:7	conditional	0. 00022	0. 00022	1. 002

Table 8 Linkage of the Relative Prices in East Asian Economies

(1) Korea CPI		Sample: 1970-1994	1994	
	Food	Clothing	Housing	Transport
Japan	1.1242	0.7998	-0.2371	1.4672
9 2 2	[2.354]	[2.8]	[-0.917]	[1.765]
The U.S	-0.0043	-0.1570	-1.5376	-0.2271
	[-0.017]	[-0.59]	[-1.44]	[-0.464]
D-₩	1.6438	1.7533	1.1997	0.7488
Ad-R2	0.4881	0.4956	0.1116	0.0210
Method	AR1, T	AR1, T	AR1, T	OLS, T

(2) Korea WPI		Sample: 1970-1994	-1994				7
	P Food	Chemical	G Machinery	E Machinery	Lumber	Plastic	lextile
	0 3003	0 2018	1	1.1091	0.6078	0.6492	1.1337
Japan	0.0200	0.60.0	1.100	, , , , ,	72 24 43	້ລວດ ເາ	ິລ ຂ∩ຄ <u>ີ</u>
	[1.7]	[0.895]	[2.745]	4.304	[3.214]	3.200	[0.000]
1	OAAOR	_0 085g	-0 1448	-0.6548	0.1867	-0.1240	-3.4542
ine U.S.	0.4100	20000	5000	ر د د د د د د د د د د د د د د د د د د د	[0 0/2]	[_0 &3Q]	[-3.58]
	2.411	[-0.314]	-0.304	[-1.5.15]	0.010	0.000	20100
2	1 4064	1 3179	1.9429	0.9245	1.7828	1.4828	0.8100
V	1,1001	2	0 0 0 0 0 0	0 085/	0 7852	0 4445	0.7670
Ad-R2	0.9467	0.0180	0.000	0.000			2
Mathod	AR1 T	AR1	AR1, T	AR1	AR1	AK-	OLO
INICHION							

			Secondary Commencer			TV
	Paper	Transpot	Non Metal	ron&Steel	Motor	- <
	0.3408	0.7940	0 3243	0.1616	1.6632	0.1372
Japan	0.0400	0.7040	0.01		[100]	[909]
	[336 c]	[2.505]	1.662	10.697	4.122	ี (0.88.0)
	[003.7]	L.0007		2000	2000	1 1011
The 115	0 1785	0.1498	0.5151	-0.33/9	-0.0993	1.1044
1.10	70.70	ביי מיי	[7227]	[-1 149]	[-0 233]	6.729
	[0.026]	[0.00]	[1,0.2]			1 01 00
J_ (≦)	1 0164	1 1996	1.0339	1.3040	1.1/10	1.8130
CVY	1.0.0			200	00000	0000
04-B0	0.8435	0.9200	0.8642	0.9/29	0.9636	0.8080
714	0.0			7	201	^ D1
Method	AR1	AR1	AK.	AXT, -	OLO, I	22
1100100					,	

Table 8 Linkage of the Relative Prices in East Asian Economies (continued)

(3) Taiwan CPI		Sample: 1970-1994	1994
	Food	Clothing	Housing
Japan	0.77738	1.21198	0.55638
	[2.204]	[4.186]	[5.069]
The U.S.	0.47627	-0.86375	-1.03809
	[2.541]	[-3.626]	[-2.461]
D-W	1.36510	1.57926	1.97880
Ad-R2	0.72808	0.97597	0.86574
Equation	OLS, T, T2	AR1, T	OLS, T, T2

(4) Taiwan WPI		Sample: 1970-1994	-1994					7	-
	P Food	Chemical	G Machinery	E Machinery	Lumber	Plastic	Textile	Paper	ransport
	1 0180	- 1	1 3687	1.7863	0.4382	0.1234	1.5809	0.5273	1.1991
Japan	1.0100	7347	[10 055]	[A 286]	[3 738]	[0 896]	[5.278]	[3.076]	[5.037]
	[4.004]	[1./04]	[10.200]	[001:1]			20007	0 5000	2 40 4 5
The U.S	0.2542	0.1415	0.2750	-0.5812	0.2767	0.8317	-2.036/	0.5302	0.4245
	[1 074]	[0.406]	[1 167]	[-1 715]	[1.962]	[6.847]	[-2.222]	[2.533]	[1.508]
	[4/2.1]	[0.100]	00041	4 4004	1 1769	1 ጙንጹን	1 4497	1 5695	0.9700
U-W	1.6040	1.1000	2.007	1.7007			2277	15.000	0 05 00
A.J. D.S	0 7362	0 7787	0.8986	0.9621	0.8116	0.8852	0.8955	0.86/1	0.8038
70	0.7001)	2 2 4	בֿרַ בַּר	יד ומא	2	O(8.7.79)	<u> </u>
Equation	AR1, T	OLS, I	OLS, I	OLS, I	AKI	77.1	OFO	000, 1, 14	010,

(5) Hong Kong CPI	CPI	Sample: 1970-1994	-1994	
	Food	Clothing	Housing	Transport
Japan	-0.2685	-0.211381	0.3891	1.5708
	[-0.767]	[-0.649]	[1.447]	[1.016]
The U.S.	0.3909	0.743015	0.362258	1.46246
	[2.072]	[3.611]	[0.323]	[1.263]
D-W	0.9432	1.47391	1.25848	1.03
Ad-R2	0.4797	0.92244	0.7848	0.3540
Method	AR1	AR1	AR1	OLS, T, T2

(6) Malaysia CPI	PI	Sample: 1970-1993	0-1993
	Food	Clothing	Transport
Japan	1.0317	0.8284	-0.1959
	[2.944]	[5.813]	[-0.599]
The U.S.	0.3496	-0.5326	0.3035
	[1.852]	[-4.272]	[0.887]
D-W	1.9085	1.8325	0.9053
Ad-R2	0.7957	0.8474	0.6239
Method	AR1, T	AR1, T	AR1

Table 8 Linkage of the Relative Prices in East Asian Economies (continued)

(7) the Philippines CPI	ines CPI	Sample: 1970-1992	-1992
	Food	Clothing	Housing
Japan	-0.35835	0.50848	0.20198
	[-0.788]	[1.993]	[0.987]
The U.S.	0.51327	-0.40056	-0.35131
	[2.389]	[-1.505]	[-0.461]
D-W	1.63290	1.27229	1.68380
Ad-R2	0.32869	0.80266	0.84925
Method	AR1	OLS, T, T2	OLS, T, T2 OLS, T, T2

(8) the Philippines WP	ines WPI	Sample: 1970-1992	-1992
	P Food	Chemical	
Japan	0.80741	0.73906	
	[2.652]	[1.352]	
The U.S.	0.34854	-0.10607	
	[1.065]	[-0.247]	
D-W	1.74935	1.57211	
Ad-R2	0.71313	0.51623	
Method	AR1	AR1	

(9) Indonesia CP	SP	Sample: 1970-1992	-1992
	Food	Clothing	Housing
Japan	-5.2754	1.2371	0.4431
	[-2.151]	[1.764]	[2.51]
The U.S.	1.3144	-2.9539	-2.2194
	[1.030]	[-4.04]	[-3.148]
D-W	2.10	0.8480	1.5933
Ad-R2	0.6109	0.9267	0.3759
Method	AR1, T	OLS, T, T2	OLS, T

(10) Indonesia WPI	WPI	Sample: 1970-1990	1990		
	G Machinery	E Machinery	Lumber	Non Metal	Iron&Steel
Japan	5.3193	1.2179	-0.0066	-0.4018	0.7503
	[5.442]	[7.04]	[-0.022]	[-1.352]	[0.739]
The U.S.	-4.8417	0.0862	0.3762	0.4632	0.8374
	[-1.89]	[0.132]	[0.995]	[1.564]	[0.93]
D-W	1.1000	1.1197	1.5612	1.1897	1.6578
Ad-R2	0.6910	0.7854	0.7077	0.7881	0.7142
Method	OLS, T	OLS, T	AR1	AR1	AR1

(11) Singapore CPI	CPI	Sample: 1970-1993	-1993	
	Food	Clothing	Housing	Transport
Japan	-0.29019	-0.08336	0.06083	1.14951
	[-0.878]	[-0.369]	[0.261]	[2.215]
The U.S.	0.51847	0.701727	1.3852	0.539954
	[2.972]	[3.489]	[1.432]	[1.161]
D-W	1.12338	1.44294	1.18186	1.9832
Ad-R2	0.82099	0.97463	0.40200	0.64674
Method	OLS, T, T2	AR1, T	AR1	AR1, T

(12) Singapore WP	WPI	Sample: 1970-1993	0-1993
	P Food	Chemical	G Machinery
Japan	1.5289	-0.3749	1.6119
	[6.423]	[-0.586]	[4.87]
The U.S.	1.0874	0.9337	0.8687
e e	[6.558]	[1.622]	[2.423]
D-W	1.9316	1.6243	1.5449
Ad-R2	0.9176	0.7309	0.8606
Method	T 'STO	AR1, T	AR1

Table 8 Linkage of the Relative Prices in East Asian Economies (continued)

(13) Thailand CPI	CPI	Sample: 1970-1993	1993	
	Food	Clothing	Housing	Transport
Japan	0.3008	0.0852	0.0449	0.1001
	[0.784]	[0.596]	[0.285]	[0.205]
The U.S.	0.1887	-0.3360	0.4092	0.9275
	[0.889]	[-2.252]	[0.629]	[1.798]
D-W	1.5140	1.2266	0.5707	1.7453
Ad-R2	0.7367	0.8696	0.7290	0.7870
Method	AR1, T	OLS, T, T2	AR1	AR1

(14) Thailand WPI	WPI	Sample: 1970-1993	-1993			
	P Food	Chemical	G Machinery	Textile	Paper	Transport
Japan	0.2585	0.0660	0.7937	1.1124	0.0894	1.1516
	[91]	[0.103]	[3.302]	[3.072]	[0.623]	[3.981]
The HS	0.1866	0.5314	0.1954	-1.5090	0.3086	-0.5720
	[1.371]	[0.983]	[0.798]	[-1.357]	[1.631]	[-1.77]
J-W	1.7872	1.3135	1.4483	1.5933	1.6755	1.7666
Ad-R2	0.3788	0.6222	0.7861	0.7481	0.6385	0.7251
Method	AR1	AR1	AR1	OLS	AR1	AR1, T
11100						

Notes of Table 8)

- t-values are in parentheses.
 "OLS" indicates that the estimation method was the ordinary square method.
 "AR1" indicates that the first-order serial correlation was corrected by the ML method.

- 4. "T" indiates that a time trend term was included in the regression.
 5. "T2" indicates that a square time trend was included in the regression.
 6. "P Food" = processed food, "G Machinery" = general machinery, and "E Machinery" = electric machinery.

Table 9 Linkage of the Relative Prices in Canada

	Food	Textile	Wood	Paper	Non-metal	Basic Metal
Tanan	0. 1661	-0.060413	0. 0633	0.1679	-0.0019	0. 4377
J cap care	[3 2]	[-1.325]	[0. 970]	[1.520]	[-0.020]	[3.683]
The II S	0. 2556	0. 163314	0. 94743	0. 251677	0.2676	0. 67166
	[11.984]	[3.464]	[21, 678]	[1.558]	[3.809]	[6.8885]
D_ W	1 8713	1. 65384	1. 39577	1.96605	1. 7506	1. 3727
Ad-R9	0.9761	0.986009	0. 6989	0.0129	0. 1170	0. 9831
me thod	AR1	ARI, T	ARI, T	AR1, T	AR1, T	ARI, T

Notes)

^{3.2.} t-values are in parentheses. "ARI" indicates that the first-order serial correlation was corrected by the ML method. "T" indiates that a time trend term was included in the regression.